DETERMINATION OF SHARE PRICE OF AGRO ALLIED FIRMS: EVIDENCE FROM NIGERIA

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Purpose: The study aims to investigate the determinant of the share price of agro-related firms listed on the Nigerian stock exchange.

Methodology: We employed regression analysis, unit root test as well as vector correction model to determine the degree of relationship between share prices and each of return on assets (ROA), earnings per share (EPS), dividend per share (DPS).

Findings: The study found that share price is majorly influenced by earnings per shares, while other variables lag it.

Practical implications: Results from the study have some possible policy implications, for instance, it is recommended that policymakers should put in place conducive market environments that will stimulate earnings from investments. Investors should on the other hand pay keen attention to information within and outside the economy when making investment decisions.

Originality/Value: This study is one of the first studies on determinants of share prices with a focus on agro-allied firms.

Keywords: Stock Market, Agro-allied Firms, Share Prices Determination, Dividend Per Share, Return on Assets, Nigeria.

INTRODUCTION

The capital market provides a platform for the sale and purchase of securities over the medium and long term (Ayoopo et al., 2015; Adelegan 2003; Fashina et al., 2018; Isola et al., 2018). Nigeria's stock market was set up to give organizations and other market practitioners a platform for raising long-term advances at negligible expenses to catalyse the nation’s assets among contentious customers.

The Nigerian stock exchange, referred to prominently as the NSE, was established in 1960 as the Lagos Stock Exchange. It changed its name to Nigerian Stock Exchange in December 1977, and currently has various branches with Lagos as its headquarters (Lawal et al., 2016; Lawal et al., 2018). Nigeria Stock Exchange gives investors the opportunities to raise funds to finance their business interests using the market mechanism. Several businesses trade on the floor of the exchange with majority trading in non-agro related activities. This raises a concern giving the fact that Nigeria is largely an agrarian economy with over 70% of her population relying on agriculture as a means of livelihood (Asaleye et al., 2018a; Lawal et al., 2017). As noted by Lawal et al., 2018, very few players on the floor of the Nigerian Stock Exchange trades in agro-allied businesses Asaleye et al., 2018b; Babajide et al., 2020; Lawal et al., 2020a). According to Lawal et al., 2019a), agro-allies firms operate in a way different from conventional firms, thus the return on investment differs in terms of seasons and volumes.

This triggers government vigorous investment in the sector with the intention of boasting this sub-sector. Despite huge government investment in the sector, the private sector response has not been encouraging. This could be due to several factors ranging from poor return on investment, low dividend payment among others. The question is what are the factors that determine the share prices of these agro-alloved firms in Nigeria?

The core of the current study is to determine the key factors influencing the prices of listed companies on the stock exchange floor of Nigerian stock exchange, with a focus on agro-allied companies. Our key goals are set out as follows:

1. To investigate whether there is a significant relationship between the return on assets (ROA) and the share price of agro-related companies.

2. To assess whether the dividend per share (DPS) has a significant relationship with the share price of agro-allied farms on the floor of the Nigeria Stock Exchange.

3. To examine the nature of the relationship between earning per share and share price of agro-allied companies.
LITERATURE REVIEW

Theoretical Framework

The theories governing the current study are: Stakeholder Theory which states that both agro-allies and stock exchange participants are key players in achieving the organization's growth and development, thus operating with stakeholders and not just shareholder mindset. Other theories are the Efficient market hypothesis theory which states that the actual price of a security will be a good estimate of its intrinsic value, and that the prices of assets fully reflect every publicly available information. (Fama, 1970) classified efficient markets into three: weak-form, semi-strong form and strong form efficient. The Capital Assets Pricing Model, which emphasizes the impact of the capital market on the aggregate economy. The hypotheses are that, by channelling money into the agricultural sector, the capital market induces development. Lawal et al., (2017b).

EMPIRICAL LITERATURE

This section provides a review of some of the existing literature on the subject matter. Lawal et al., (2018) examined the impact of returns on assets, return on equity, and debt to equity on stock return in wholesale and retail trade companies listed on the Indonesia stock exchange. The study employed multiple linear regression and noted that return on assets has a negative effect on stock returns, while both the return on equity and debt to equity ratio have a positive impact on stock return for the studied economy. These findings was supported by the study of Ilmiyono (2019) and Idawati and Wahyudi (2015) who both noted that the impact of return on the asset on stock market prices is not always positive but time-varying.

Mogonta and Pandowo (2016) employed multiple regression to analyze dataset on return on asset (RoA), return on equity (RoE), and earnings per share (EPS) for seven (7) mining firms listed on the Indonesian stock market based on data sourced from 2011 to 2015. The study noted that the three variables have a varying degree of influence on stock market prices. This position was supported by the findings of Utami and Darmawan (2019) who employed a purposive sampling method to analyse data sourced from 53 companies on the Indonesia stock exchange from the period 2012 to 2016.

Pandey (2019) examined the effect of both macroeconomic variables proxy by interest rate, exchange rate, tax rate, inflation rate; and microeconomic variables like the return on asset (RoA), return on equity (RoE) and earnings per share (EPS) on share price determinants in India pharmaceutical firms listed on the Indian stock exchange and some selected developed economies comprises of the USA, UK, Germany, France, and Switzerland. The study noted that pharmaceutical firms share prices of the studied economies are largely influenced by macroeconomic variables with microeconomic variables like return on asset, return on equity having little or no impact on stock price movement.

Sharif, Purohit, and Pillai (2015) examined the determinants of stock prices movement for Bahrain based on data sourced from the year 2006 to 2010. The study employed pooled OLS regression with robust standard errors, fixed effects and random effect techniques and observed that return on equity, book value per share, dividend per share, dividend yield, price-earnings and firm size are significant determinants of share price in the Bahrain stock market while debt to assets and earnings per share have no impact on stock prices.

For the Pakistania economy, Hunira et al., (2014) examined the determinants of stock prices of four non-financial sectors – sugar, chemical, food and personal care, and energy – using the OLS to analyse data sourced from 2006 to 2011. The study noted that dividend is the prime determinant of share prices. This result was in line with the study of Zainudin, Mahdzan, and Yet (2017) for the Malaysian stock market, who noted that dividend policy is the major determinants of share price movement.

For the Turkish economy, Kassouri and Altintas (2020) examined the impact of exchange rate volatility on stock market price determination. The study employed complex asymmetric models and a non-linear framework to analyze monthly data sourced from January 2003 to December 2018. The study noted that the impact of the exchange rate on stock market pricing is better revealed from a non-linear asymmetric model.

Aatola, Ollikainen, and Toppinen (2013) examined the price determination of the European Union Emission Allowance (EUA) of the European Union Emission Trading Scheme (EU ETS). The study postulated an uncertain permit price and risk-averse firms that can hedge in a forward market system. The study employed several econometrics models and a battery of unit root techniques to analyze a daily dataset sourced from 2015 to 2010 in the EU ETS markets. The study noted that evidence abound to state that a significant relationship exists among the German electricity prices, coal prices, stock market prices and the market fundamentals and the EUA prices. The study also noted that the EUA forward price depends largely on the fundamentals.
In a related development, Hintermann, Peterson, and Rickels (2016) examined the links between price and market behavior of EU ETS focusing on the second phase period. The study noted that allowance prices are relatively lower for the studied economies (see also Zhang and Sun, 2016).

Ariff, Chung, and Shamsher (2012) examined the impact of interest rate, money supply, and liquidity effect on share price determination. The study employed a battery of simultaneous equations to analyze quarterly data sourced from 1960 to 2011 and noted that liquidity effect and money supply impact positively in determining asset prices for the studied economy. The study noted that the money supply is endogenously influenced, following a post-Keynesian economic model.

For China, Geertsema and Lu, (2019) employed simultaneous equations to examine the existence of bi-directional causality between Initial Public Offer (IPO) initial returns and oversubscriptions ratios in order to determine factors that induce price formation. The study scope ranges from 1996 to 2015. The study noted that a unidirectional relationship exists from oversubscription ratios to IPO initial returns, and that price is largely determined by market forces in China.

For electrolytic related prices, Shi, Wang, Feng, and Wu, (2018) employed a structural demand-supply model to examine what influences price determination. The study noted that the prices of electricity-related stock are largely influenced by the transmission effect of the equilibrium price of the electrolytic mechanism.

For the Australian dollar market, Su and Zhang (2018) examined factors that motivate price discovery based on intraday trading data sourced from January 1999 to December 2013. The study employed a battery of econometric models and observed that macroeconomic news, order flows, and market state variables such as returns, volatility, trade volume and bid/ask spread are key determinants of price discovery in the studied economy.

METHODOLOGY

The data for this study were sourced from the publication of the Central Bank of Nigeria (various issues) and the Nigerian Stock Exchange (various issues). It comprises annual data sourced from 2006 to 2018. The period is strategic as its starting date precedes the 2007/2008 global financial crisis and the end date is after the recovery from the economic recession of 2016Q3 to 2017Q2 witnessed in Nigeria. We used regression analysis, unit root test as well as vector correction model to analyse our results. We first estimated their individual stationary levels by using the Augmented Dickey Fuller test which helps to test the time series data for unit root existence. We then employed a simple regression model and variance decomposition estimates to examine the relationship among the variables.

The specification for the regression model is as follows.

\[ Y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_3 + \beta_4 x_4 + \mu \]

Where \( y \) is an independent variable while \( x_1-4 \) are the independent variables, \( Y \) = the share price of listed agro-allied firms represent by return on asset (ROA), \( \beta \) = constant term, \( X_1 \) = share price of agro-allied firms (SHP), \( X_2 \) = return on asset (ROA), \( X_3 \) = dividend per share (DPS), \( X_4 \) = earnings per shares (EPS), \( \mu \) = the error terms.

RESULTS/FINDINGS

Unit Root Test

The Augmented Dickey-Fuller test was performed for the variables used to check for both stationary and non-stationary results.

<table>
<thead>
<tr>
<th></th>
<th>T-stat</th>
<th>Probability</th>
<th>Significance level</th>
<th>Integration order</th>
</tr>
</thead>
<tbody>
<tr>
<td>SHP</td>
<td>-3.178569</td>
<td>0.0017</td>
<td>At level</td>
<td>I(0)</td>
</tr>
<tr>
<td>ROA</td>
<td>-4.013928</td>
<td>0.0001</td>
<td>At level</td>
<td>I(0)</td>
</tr>
<tr>
<td>DPS</td>
<td>-3.797985</td>
<td>0.0002</td>
<td>At level</td>
<td>I(0)</td>
</tr>
<tr>
<td>EPS</td>
<td>-2.879881</td>
<td>0.0042</td>
<td>At level</td>
<td>I(0)</td>
</tr>
</tbody>
</table>

Source: Author’s computation with EViews 10

Table 1 above provides a description of the results of unit root tests obtained for the various variables at the point. The Augmented Dickey Fuller test was conducted on all variables and the results showed that the variables were stationary at the level of 1%, 5% and 10% of the critical value because their T statistics at the levels were higher than the critical value of 1%, 5% and 10%.
Linear Regression Analysis (Ordinary Least Square)

Table 2 summarizes the findings of the linear regression analysis performed to determine if a significant relationship exists between the independent variables and the dependent variable. The condition for the OLS regression is that each variable's likelihood value should be less than 0.05 at a significance level of 5 percent, for each variable, the T-Statistics value should be greater than 2 for each variable. DPS meets the three criteria and thus shows that the relationship between DPS and SHP is significant. On the other hand, ROA and EPS do not meet the criteria, and hence the null hypothesis is accepted for their case.

Table 2: Linear Regression Analysis (Ordinary Least Square)

<table>
<thead>
<tr>
<th>Variables</th>
<th>Coefficient</th>
<th>Standard error</th>
<th>T-stat</th>
<th>Probability</th>
</tr>
</thead>
<tbody>
<tr>
<td>SHP</td>
<td>1.564309</td>
<td>1.564007</td>
<td>1.96003</td>
<td>0.07455</td>
</tr>
<tr>
<td>ROA</td>
<td>1.486968</td>
<td>1.800215</td>
<td>0.825995</td>
<td>0.08104</td>
</tr>
<tr>
<td>DPS</td>
<td>29.07226</td>
<td>1.511979</td>
<td>19.22795</td>
<td>0.0000</td>
</tr>
<tr>
<td>EPS</td>
<td>0.001286</td>
<td>0.102020</td>
<td>0.012601</td>
<td>0.0900</td>
</tr>
<tr>
<td>C</td>
<td>-0.866729</td>
<td>10.15467</td>
<td>-0.853472</td>
<td>0.0650</td>
</tr>
</tbody>
</table>

Source: Author’s computation 2020

Variance Decomposition

Thanks to their own shocks and the shocks of the exogenous variables ROA, DPS and EPS, the variance decomposition would be used to evaluate the magnitude of the changes within the dependent variable SHP.

Table 3: Variance Decomposition of ROA

<table>
<thead>
<tr>
<th>Period</th>
<th>S.E.</th>
<th>SHP</th>
<th>ROA</th>
<th>DPS</th>
<th>EPS</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>2.872036</td>
<td>0.001106</td>
<td>99.99889</td>
<td>0.000000</td>
<td>0.000000</td>
</tr>
<tr>
<td>2</td>
<td>3.605150</td>
<td>0.009369</td>
<td>99.74248</td>
<td>0.248080</td>
<td>7.06E-05</td>
</tr>
<tr>
<td>3</td>
<td>4.042979</td>
<td>0.012711</td>
<td>99.44690</td>
<td>0.419751</td>
<td>0.120640</td>
</tr>
<tr>
<td>4</td>
<td>4.315271</td>
<td>0.01478</td>
<td>99.16848</td>
<td>0.461693</td>
<td>0.358350</td>
</tr>
<tr>
<td>5</td>
<td>4.493649</td>
<td>0.013820</td>
<td>98.89496</td>
<td>0.480487</td>
<td>0.610735</td>
</tr>
</tbody>
</table>

Source: Author’s computation with Eviews 10

Table 4 above shows the decomposition of ROA variances. It can be deduced from the table that the variability in ROA's forecast error shock is explained by about 99.9 percent variance in the first half. The variance in ROA's forecast error shock is explained by around 0.24 percent of the variability in the variable DPS in the second period. The variation in ROA's forecast error shock is explained by about 0.12 percent of the variation in EPS in the third period. The variation in the ROA's forecast error shock is explained by about 0.013 percent of the variation in SHP in the fourth period.

Table 4: Variance Decomposition of DPS

<table>
<thead>
<tr>
<th>Period</th>
<th>S.E.</th>
<th>SHP</th>
<th>ROA</th>
<th>DPS</th>
<th>EPS</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>3.183229</td>
<td>29.90621</td>
<td>0.043260</td>
<td>70.05053</td>
<td>0.000000</td>
</tr>
<tr>
<td>2</td>
<td>4.178576</td>
<td>56.65610</td>
<td>0.118840</td>
<td>43.15398</td>
<td>0.071078</td>
</tr>
<tr>
<td>3</td>
<td>4.694356</td>
<td>64.14157</td>
<td>0.095862</td>
<td>35.70279</td>
<td>0.059777</td>
</tr>
<tr>
<td>4</td>
<td>5.101583</td>
<td>62.88990</td>
<td>0.089497</td>
<td>36.96146</td>
<td>0.059140</td>
</tr>
<tr>
<td>5</td>
<td>5.379319</td>
<td>62.59849</td>
<td>0.096850</td>
<td>37.23865</td>
<td>0.066008</td>
</tr>
</tbody>
</table>

Source: Author’s computation with EViews 10

Table 4 above shows the decomposition of DPS variance. It can be deduced from the table that the variability in DPS' forecast error shock is clarified in itself by about 70.05 percent variance in the first period. The variability in the DPS forecast error shock is explained by around 0.12 percent of the variance in the ROA variable in the second period. The variance in DPS' forecast error shock is explained by around 0.06 percent of the variability in EPS in the third period. The variance in the DPS' forecast error shock is explained by about 62.6 percent of the variability in SHP in the fourth period.
Table 5: Variance Decomposition of EPS

<table>
<thead>
<tr>
<th>Periods</th>
<th>S. E.</th>
<th>SHP</th>
<th>ROA</th>
<th>DPS</th>
<th>EPS</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>64.12466</td>
<td>0.003791</td>
<td>1.400333</td>
<td>0.484984</td>
<td>98.11089</td>
</tr>
<tr>
<td>2</td>
<td>79.81568</td>
<td>0.059526</td>
<td>1.301045</td>
<td>0.862552</td>
<td>97.77688</td>
</tr>
<tr>
<td>3</td>
<td>84.00364</td>
<td>0.057838</td>
<td>1.176663</td>
<td>1.118354</td>
<td>97.64714</td>
</tr>
<tr>
<td>4</td>
<td>84.97183</td>
<td>0.074419</td>
<td>1.201364</td>
<td>1.142369</td>
<td>97.58185</td>
</tr>
<tr>
<td>5</td>
<td>85.22952</td>
<td>0.104712</td>
<td>1.295091</td>
<td>1.150287</td>
<td>97.44991</td>
</tr>
</tbody>
</table>

Source: Author’s computation with EViews 10

Table 5 above shows the decomposition of the EPS variance. It can be deduced from the table that the variability in EPS forecast error shock is explained in itself by about 98.1 percent variance in the first half. The variability in the projected error shock of EPS is clarified in the second period by about 1.30 percent of the variance in the variable ROA. The variation in EPS forecast error shock is explained by about 1.11 percent of the variance in DPS in the third period. The variation in the EPS forecast error shock is explained by around 0.10 percent of the variation in SHP in the fourth period.

It can be noticed from the diagram above that all the blue dots are inside the circle. That means the model is stable. Model stability means which economic predictions can be made on a model basis.

Overall, the results show that the share price of agro-allied firms in Nigeria is largely influenced by earnings per share. This is in line with existing studies like Mogonta and Pandowo (2016) for the mining firms listed on the Indonesia stock exchange. The results also correspond with the findings of Hunira et al, (2014) for Pakistan. However, the result contradicts the findings of Saragih et al, (2018) who documented a negative relationship between share prices and return on assets. The results were also at variance with the findings of Pandey (2019), who noted that macroeconomic fundamentals like inflation rate, exchange rate, interest rate, tax rate are the key determinants of share prices in India and the Western economies. Our results also differ from the findings of Sharif et al, (2015), who noted that earnings per share have no impact on share price determinations for the Bahrain economy. Unlike the findings of Saragih (2018), who documented the existence of a negative relationship between RoA and share prices, our results observed the existence of the positive but insignificant relationship between RoA and share prices, our results was similar to the findings of Ilmiyono and Idawati and Wahyudi (2015) who both noted that the impact of return on the asset on stock market prices time-varying. The differences between our results and some of the existing studies could be as a result of the size and scope of the variables employed, the nature of the industry studied, among others. For instance, Saragih (2018), employed fewer dataset, Mogonta and Pandowo (2016) focused on mining sub-sector with have almost the same characteristics like the agro-allied firms in our case.

CONCLUSION

This research work analysed the share price of agro-allied and Non agro allied companies listed on the stock exchange market in Nigeria. This study is based on annual variable results, a total of 12 years, starting from 2006 to 2018. This research work analysed the share price of agro allied and Non agro allied companies listed on the stock exchange market in Nigeria. In the analysis, the ordinary least square regression approach was used for the data collected prior to the stationary test and after the stationary test, unit root test was also performed to know at what level the variable is stationary, descriptive statistics were collected to evaluate the mean and standard deviation, correlation testing was also conducted to evaluate the relation between dependent and independent variables. The hypotheses were tested to determine the level or percentage of the dependent-independent variable relation. In this study, all of the independent variables affect the financial profitability of agro allied and Non agro allied companies.

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AUTHORS CONTRIBUTION

Author One did the conceptualization and the first and final drafts, Authors Two and Three did the Methodology and Results while Authors Four did the literature review.
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